# Online Appendix 

# "Liquidity Constraints, Spillovers, and Entrepreneurship: Direct and Indirect Effects of a Cash Transfer Program' 

May 4, 2018

## A IV with a Cluster-Level Instrument

Proposition A.1. Let $z_{i v t}$ be an instrumental variable. If the period-cluster conditional variance of $z_{i v t}$ is zero, $\operatorname{Var}\left(z_{i v t} \mid v, t\right)=0$, then the IV estimator for $\tau$ in equation (4.2) is equivalent to the IV estimator for $\tau$ in the following equation:

$$
\begin{equation*}
y_{i v t}=\beta_{0}+\tau d_{i v t}+\mu_{v}+\mu_{t}+u_{i v t} . \tag{A.1}
\end{equation*}
$$

Proof. Let $y_{i v t}^{*}, d_{i v t}^{*}$, and $\bar{d}_{v t}^{*}$ be cluster-period mean-centered versions of $y_{i v t}, d_{i v t}$, and $\bar{d}_{v t}$, respectively.

Suppose equation (4.1) is the true equation, but we instead estimate the following model:

$$
\begin{equation*}
y_{i v t}=\beta_{0}+\beta_{1} d_{i v t}+\mu_{v}+\mu_{t}+u_{i v t}, \tag{A.2}
\end{equation*}
$$

in which $\bar{d}_{v t}$ is omitted.
Let $z_{v t}$ be an instrumental variable such that $\operatorname{Var}\left(z_{v t} \mid v, t\right)=0$. Then the (within-group) IV estimator for $\beta_{1}$ in equation (A.2) is:

$$
\begin{aligned}
\widehat{\beta}_{1}^{I V} & =\frac{\sum_{i v t} z_{v t}^{*} y_{i v t}^{*}}{\sum_{i v t} z_{v t}^{*} d_{i v t}^{*}} \\
& =\frac{\sum_{i v t} z_{v t}^{*} y_{i v t}^{*}}{\sum_{v t} z_{t v}^{*} \sum_{i} d_{i v t}^{*}} \\
& =\frac{\sum_{i v t} z_{v t}^{*} y_{i v t}^{*}}{\sum_{i v t} z_{v t}^{*} \overline{\bar{d}}_{v t}^{*}}=\widehat{\tau}^{I V} .
\end{aligned}
$$

Thus the formula is exactly the same as if we estimate equation (4.2) using $z_{v t}$ as an instrumental variable. Using similar steps as in Proposition 4.1, we can show that $\widehat{\tau}^{I V}$, as well as $\widehat{\beta}_{1}^{I V}$, is a consistent estimator for the overall effect, $\left(\beta_{1}+\beta_{2}\right)$.
Table B1: Poverty Headcount and Program Coverage

|  | 2001 |  |  | 2004 |  |  | 2006 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | Urban | Rural | Total | Urban | Rural | Total | Urban | Rural |
| Poverty headcount | 0.301 | 0.250 | 0.579 | 0.285 | 0.241 | 0.534 | 0.225 | 0.183 | 0.466 |
| Program coverage | 0.064 | 0.044 | 0.174 | 0.178 | 0.146 | 0.360 | 0.227 | 0.188 | 0.450 |
| Share of benefits |  | 0.599 | 0.401 |  | 0.686 | 0.314 |  | 0.708 | 0.292 |
| Number of obs. | 368,605 | 316,793 | 51,812 | 378,658 | 326,322 | 52,336 | 389,807 | 336,502 | 53,305 |

Estimates are obtained using PNAD. 'Poverty headcount' is measured by the proportion of people with household per capita income below the poverty line (half the total amount of transfers going to either urban or rural areas and the total amount of transfers distributed by CCT programs in the country.

Table B2: Number of Observations per Municipality

|  |  | Std. |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Mean | Dev. | Min. | Max. | Number of <br> municipalities |  |
| 2001 |  |  |  |  |  |
| Number of households | 128.1 | 290.4 | 19 | 3,505 | 796 |
| Sample size | 52.4 | 128.1 | 5 | 1,571 | 796 |
| $\mathbf{2 0 0 4}$ |  |  |  |  |  |
| Number of households | 136.8 | 305.1 | 23 | 3,575 | 796 |
| Sample size | 54.3 | 131.8 | 5 | 1,751 | 796 |
| 2006 |  |  |  |  |  |
| Number of households | 143.8 | 322.7 | 28 | 3,884 | 796 |
| Sample size | 56.4 | 136.1 | 5 | 1,753 | 796 |

The sample comprises men aged between 25 and 45 years old, with no college degree, and living in urban areas. This sample also excludes public servants and employers with more than five employees.

Table B3: Overall Effect of Cash Transfers on Different Types of Business

|  | Decision of being a small entrepreneur in |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Services |  | Sales |  | Manufacturing |  |
|  | FE <br> (1) | $\begin{aligned} & \text { IV } \\ & (2) \\ & \hline \end{aligned}$ | $\begin{aligned} & \mathrm{FE} \\ & (3) \\ & \hline \end{aligned}$ | IV <br> (4) | $\begin{aligned} & \mathrm{FE} \\ & (5) \\ & \hline \end{aligned}$ | IV <br> (6) |
| program coverage, $\bar{d}$ | $\begin{aligned} & 0.040^{* * *} \\ & (0.012) \end{aligned}$ | $\begin{aligned} & 0.056^{* * *} \\ & (0.017) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.009) \end{aligned}$ | $\begin{aligned} & 0.015 \\ & (0.013) \end{aligned}$ | $\begin{aligned} & 0.010 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.006 \\ & (0.011) \end{aligned}$ |
| age (x10) | $\begin{aligned} & 0.031^{* * *} \\ & (0.012) \end{aligned}$ | $\begin{aligned} & 0.031^{* *} \\ & (0.012) \end{aligned}$ | $\begin{aligned} & 0.023^{* *} \\ & (0.012) \end{aligned}$ | $\begin{aligned} & 0.023^{* *} \\ & (0.012) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.010) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.010) \end{aligned}$ |
| squared age (x100) | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.002) \end{aligned}$ |
| white | $\begin{aligned} & 0.016^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.015^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.015^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.015^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* * *} \\ & (0.001) \end{aligned}$ |
| black | $\begin{aligned} & -0.006^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.006^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.005^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.005^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.005^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.005^{* * *} \\ & (0.001) \end{aligned}$ |
| married | $\begin{aligned} & 0.000 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.012^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.012^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* * *} \\ & (0.001) \end{aligned}$ |
| elementary education | $\begin{aligned} & 0.011^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.011^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.011^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.011^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.008^{* * *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.008^{* * *} \\ & (0.001) \end{aligned}$ |
| primary education | $\begin{aligned} & 0.012^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.012^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.015^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.015^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.003^{*} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.003^{*} \\ & (0.002) \end{aligned}$ |
| high school | $\begin{aligned} & 0.022^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.022^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.013^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.013^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ |
| log of population | $\begin{aligned} & -0.010 \\ & (0.011) \end{aligned}$ | $\begin{aligned} & -0.011 \\ & (0.011) \end{aligned}$ | $\begin{aligned} & -0.015 \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.015 \\ & (0.009) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.008) \end{aligned}$ |
| year $=2001$ | $\begin{aligned} & 0.020^{* * *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & 0.023^{* * *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.009^{* * *} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.008^{* * *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.003 \\ & (0.002) \end{aligned}$ |
| year $=2004$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.001) \end{aligned}$ |
| Municipality Fixed-Effects | Yes | Yes | Yes | Yes | Yes | Yes |
| Number of observations | 112,117 | 112,117 | 112,117 | 112,117 | 112,117 | 112,117 |

${ }^{* * *},{ }^{* *}, *$ represent statistical significant at the $1 \%, 5 \%$ and $10 \%$ levels, respectively. Standard errors in parentheses are clustered by municipality. Sample includes only men with high school diploma or less. FE columns present the fixed-effect regressions obtained using the within-group method. IV columns present the fixed-effect, InstrumentalVariable regressions with 'program coverage' instrumented by the interactions between municipal quotas and year dummies.

Table B4: Indirect and Direct Effects on Entrepreneurship, With and Without Children

|  | Decision of being a small entrepreneur |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
|  | Without children |  | With children |  |
|  | FE | IV | FE | IV |
|  | (1) | (2) | (3) | (4) |
| program coverage, $\bar{d}$ | 0.028 | 0.044 | 0.090*** | 0.112*** |
|  | (0.021) | (0.029) | (0.025) | (0.030) |
| individual benefit, $d$ | $-0.013^{* * *}$ | $-0.013^{* * *}$ | -0.050*** | -0.058*** |
|  | (0.005) | (0.005) | (0.009) | (0.012) |
| age (x10) | 0.071** | 0.071** | 0.062** | 0.062** |
|  | (0.029) | (0.029) | (0.028) | (0.028) |
| squared age (x100) | -0.004 | -0.004 | -0.002 | -0.002 |
|  | (0.004) | (0.004) | (0.004) | (0.004) |
| white | 0.031*** | 0.031*** | $0.037^{* * *}$ | $0.037^{* * *}$ |
|  | (0.002) | (0.002) | (0.003) | (0.003) |
| black | -0.010** | -0.010** | $-0.021^{* * *}$ | $-0.021^{* * *}$ |
|  | (0.004) | (0.004) | (0.004) | (0.004) |
| married | 0.027*** | 0.027*** | 0.028*** | 0.028*** |
|  | (0.002) | (0.002) | (0.003) | (0.003) |
| elementary education | 0.027*** | 0.027*** | 0.029*** | 0.028*** |
|  | (0.003) | (0.003) | (0.003) | (0.003) |
| primary education | 0.027*** | 0.027*** | 0.029*** | 0.029*** |
|  | (0.003) | (0.003) | (0.004) | (0.004) |
| high school | 0.029*** | 0.029*** | 0.031*** | 0.031*** |
|  | (0.003) | (0.003) | (0.004) | (0.004) |
| log of population | 0.002 | 0.001 | $-0.064^{* * *}$ | -0.065*** |
|  | (0.022) | (0.022) | (0.023) | (0.023) |
| year $=2001$ | 0.004 | 0.006 | 0.001 | 0.005 |
|  | (0.005) | (0.006) | (0.006) | (0.007) |
| year $=2004$ | -0.002 | -0.001 | -0.002 | -0.001 |
|  | (0.003) | (0.003) | (0.004) | (0.004) |
| Municipality Fixed-Effects | Yes | Yes | Yes | Yes |
| N. of obs. - all sample | 63,348 | 63,348 | 65698 | 65698 |
| N. of obs. $-d=0$ | 60,630 | 60,630 | 52,458 | 52,458 |

${ }^{* * *},{ }^{* *}, *$ represent statistical significant at the $1 \%, 5 \%$ and $10 \%$ levels, respectively. Sample includes only men with high school diploma or less. Standard errors in parentheses are clustered by municipality. All coefficients are estimated using Seemingly Unrelated Regressions (SUR). The indirect effect (program coverage) is estimated using the sample of non-participants, whereas the direct effect (individual benefit) is estimated using all sample and bias corrected according to Lemma 4.1. Columns (1) and (2) present the estimates of effects on individuals without children in their household. Columns (3) and (4) present the estimates of effects on individuals living with children under 15 years old. The FE column shows the fixed-effect regression obtained using the within-group method. The IV column shows fixed-effect, Instrumental-Variable regression with 'program coverage' instrumented by the interactions between municipal quotas and year dummies.

Table B5: Indirect and Direct Effects on Occupational Choices, With and Without High School

| Panel A: Individuals without High-School Diploma |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Fixed-Effect Model |  |  |  |  |
|  | Entrep. | Jobless | Formal employee | Informal employee | Informal self-emp. |
| program coverage, $\bar{d}$ | 0.05*** | 0.023 | 0.04 | -0.087** | -0.025 |
|  | (0.016) | (0.026) | (0.035) | (0.034) | (0.035) |
| individual benefit, $d$ | $-0.038^{* * *}$ | 0.035*** | $-0.044^{* * *}$ | 0.026* | 0.021 |
|  | (0.006) | (0.011) | (0.014) | (0.014) | (0.015) |
| Municipality Fixed-Effe Year dummies <br> Individual covariates <br> N. of obs. - all sample <br> N. of obs. $-d=0$ | Yes | Yes | Yes | Yes | Yes |
|  | Yes | Yes | Yes | Yes | Yes |
|  | Yes | Yes | Yes | Yes | Yes |
|  | 90,648 | 90,648 | 90,648 | 90,648 | 90,648 |
|  | 76,566 | 76,566 | 76,566 | 76,566 | 76,566 |
|  | Instrumental-Variable Model |  |  |  |  |
|  | Entrep. | Jobless | Formal employee | Informal employee | Informal self-emp. |
| program coverage, $\bar{d}$ | 0.064*** | 0.055 | -0.007 | -0.109*** | -0.004 |
|  | (0.022) | (0.034) | (0.047) | (0.041) | (0.045) |
| individual benefit, $d$ | $-0.043^{* * *}$ | $0.043^{* * *}$ | $-0.051^{* * *}$ | $0.01$ | $0.042^{* *}$ |
|  | (0.007) | $(0.016)$ | $(0.018)$ | (0.019) | $(0.021)$ |
| Municipality Fixed-Effects Year dummies | Yes | Yes | Yes | Yes | Yes |
|  | Yes | Yes | Yes | Yes | Yes |
| Individual covariates <br> N. of obs. - all sample <br> N. of obs. $-d=0$ | Yes | Yes | Yes | Yes | Yes |
|  | 90,648 | 90,648 | 90,648 | 90,648 | 90,648 |
|  | 76,566 | 76,566 | 76,566 | 76,566 | 76,566 |
| Panel B: Individuals with High-School Diploma |  |  |  |  |  |
|  | Fixed-Effect Model |  |  |  |  |
|  | Entrep. | Jobless | Formal employee | Informal employee | Informal self-emp. |
| program coverage, $\bar{d}$ | $0.086^{*}$ | $0.017$ | $-0.037$ | $-0.014$ | $-0.052$ |
|  | (0.047) | (0.048) | $(0.073)$ | (0.035) | $(0.039)$ |
| individual benefit, $d$ | $-0.039^{* * *}$ | 0.03** | -0.053** | 0.029** | 0.034** |
|  | (0.011) | (0.013) | (0.021) | (0.014) | (0.015) |
| Municipality Fixed-Effe <br> Year dummies <br> Individual covariates <br> N. of obs. - all sample <br> N. of obs. $-d=0$ | Yes | Yes | Yes | Yes | Yes |
|  | Yes | Yes | Yes | Yes | Yes |
|  | Yes | Yes | Yes | Yes | Yes |
|  | 38,398 | 38,398 | 38,398 | 38,398 | 38,398 |
|  | 36,522 | 36,522 | 36,522 | 36,522 | 36,522 |
|  | Instrumental-Variable Model |  |  |  |  |
|  | Entrep. | Jobless | Formal employee | Informal employee | Informal self-emp. |
| program coverage, $\bar{d}$ | $0.103^{*}$ | $0.003$ |  |  | $-0.061$ |
|  | (0.056) | $(0.054)$ | (0.082) | (0.045) | $(0.051)$ |
| individual benefit, $d$ | $-0.038^{* * *}$ | 0.026 | -0.04* | 0.011 | 0.041** |
|  | (0.012) | (0.017) | (0.024) | (0.017) | (0.018) |
| Municipality Fixed-Effects | Yes | Yes | Yes | Yes | Yes |
| Year dummies | Yes | Yes | Yes | Yes | Yes |
| Individual covariates | Yes | Yes | Yes | Yes | Yes |
| N. of obs. - all sample | 38,398 | 38,398 | 38,398 | 38,398 | 38,398 |
| N . of obs. $-d=0$ | 36,522 | 36,522 | 36,522 | 36,522 | 36,522 |

***, **, * represent statistical significant at the $1 \%, 5 \%$ and $10 \%$ levels, respectively. Standard errors in parentheses are clustered by municipality. All coefficients are estimated using Seemingly Unrelated Regressions (SUR). The indirect effect (program coverage) is estimated using the sample of non-participants, whereas the direct effect (individual benefit) is estimated using all sample and bias corrected according to Lemma 4.1. Fixed-Effect models are estimated using the within-group method. In the Instrumental-Variable models, 'program coverage' is instrumented by the interactions between municipal quotas and year dummies.

